



Statistics for Astronomers: Lecture 6, 2020.10.14

Prof. Sundar Srinivasan

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Frequentist statistical inference:
Parametric (specify model, compute likelihood) vs. nonparametric (performed on rank-ordered data).
Estimation (point/interval) or hypothesis testing.
Bayesian vs frequentist inference.
Statistics and their desired properties.
Estimators, estimates. Bias-variance tradeoff.
Point estimates: likelihood.



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A method of point estimation.

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"The likelihood that [the parameters] should have [an assigned set of values] is proportional to the probability that if this were so, the totality of observation should be that observed."

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1D example: N = 10 coin tosses result in X = 8 heads. Estimate P(H).

$$\mathscr{L}(heta)=\mathsf{P}(\mathsf{X}=\mathsf{8},\mathsf{N}=\mathsf{10}\mid heta)={10 \choose 8}\ heta^8\ (1- heta)^2,$$
 with $0< heta<\mathsf{1}.$



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$$\frac{\partial}{\partial \theta} \ln \mathscr{L}(\theta) = \frac{8}{\theta} - \frac{2}{1-\theta}; \qquad \qquad \text{vanishes at } \theta = \hat{\theta}_{\text{\tiny MLE}} \Longrightarrow \hat{\theta}_{\text{\tiny MLE}} = \textbf{0.8}.$$



 $\vec{\theta} = (\mu, \sigma^2)$. *N* observations $X_i (i = 1, \cdots, N) \sim \mathcal{N}(\mu, \sigma^2)$.



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$$\begin{split} & \frac{\partial \ell}{\partial \mu} = \sum_{i=1}^{N} \left(\frac{x_i - \mu}{\sigma^2} \right) \\ & \text{@ MLE:} \sum_{i=1}^{N} \left(\frac{x_i - \hat{\mu}}{\widehat{\sigma^2}} \right) = \mathbf{0} \end{split}$$

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@ MLE: $\sum_{i=1}^{N} \left(\frac{x_i - \hat{\mu}}{\sigma^2} \right) = 0$
 $\implies \hat{\mu} = \frac{1}{N} \sum_{i=1}^{N} x_i \equiv \bar{x}.$

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$$\mu$$
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$$\frac{\partial \ell}{\partial \sigma} = \frac{1}{\sigma} \left(-N + \sum_{i=1}^{N} \frac{(x_i - \hat{\mu})^2}{\sigma^2} \right)$$

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 $\widehat{\sigma^2} = \frac{1}{2} \sum_{i=1}^{N} (x_i - \hat{\mu})^2 - \frac{1}{2} \sum_{i=1}^{N}$

$$\implies \widehat{\sigma^2} = \frac{1}{N} \sum_{i=1}^N (x_i - \hat{\mu})^2 = \frac{1}{N} \sum_{i=1}^N (x_i - \bar{x})^2$$

MLE of σ^2 is the (biased) sample variance!



Coin-toss problem: assume that the true θ is $\theta_0 = 0.8$, unknown to observer. Each round of ten tosses: different value of $\hat{\theta}_{\text{MLE}}$ (e.g., [0.9, 0.7, 0.9, 0.8, 1., 0.9, 1., 0.9, 0.9, 0.8]). With finite # experiments, not enough to just quote $\hat{\theta}_{\text{MLE}}$. What is the variance on the MLE?



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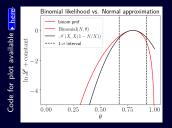
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 $\ln \mathscr{L}$ "regular" if we can ignore higher-order terms.



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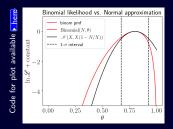


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Curvature defined as the negative second derivative of $\ln \mathscr{L}$ at location of maximum: $I(\theta) \equiv -\frac{\partial^2}{\partial \theta^2} \log \mathscr{L}$ (1-D) $I_{ij}(\vec{\theta}) \equiv -\frac{\partial}{\partial \theta_i} \frac{\partial}{\partial \theta_j} \log \mathscr{L}$ (N-D) Fisher information matrix. Large curvature near $\vec{\theta_0}$: less uncertainty (more information) about location of maximum.



N-D Taylor Expansion:
$$\ln \left[\frac{\mathscr{L}(\vec{\theta})}{\mathscr{L}(\vec{\theta}_0)}\right] = -\frac{1}{2} (\vec{\theta} - \vec{\theta}_0) \underbrace{\left[-\frac{\partial}{\partial \vec{\theta}} \frac{\partial}{\partial \vec{\theta}} \ln \mathscr{L}(\theta)\right]}_{\theta_0} (\vec{\theta} - \vec{\theta}_0)^{\mathrm{T}}$$



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To compare with the true value, define:

Average/Expected Fisher information:
$$\mathcal{I}(\vec{\theta}) \equiv \mathbb{E}[I(\vec{\theta})] = \mathbb{E}\left[-\frac{\partial}{\partial \vec{\theta}}\frac{\partial}{\partial \vec{\theta}}\log \mathscr{L}\right].$$



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The inverse of the Expected Fisher matrix is the covariance matrix of the parameters:

$$\Sigma(\vec{\theta}) = \mathcal{I}^{-1}(\vec{\theta})$$



Experiment: Ten coin tosses with unknown probability θ of obtaining a head.

 $\mathscr{L}(\theta) \propto \theta^X (1-\theta)^{(N-X)}$, and $\mathbb{E}[X] = N\theta$.



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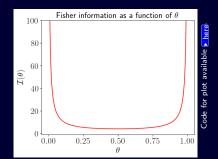
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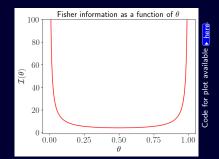
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Variance of
$$\operatorname{Binomial}(N, \theta) = N \; \theta \; (1 - \theta).$$
$$= \frac{1}{\mathcal{I}(\theta)} \; \text{in this case!}$$





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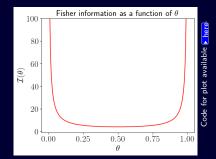
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Is this always true? Cramér-Rao Lower Bound.



If $\vec{\mathbf{T}}(X)$ is an unbiased estimator of a function $\vec{\mathbf{g}}(\vec{\theta})$ of the parameters $\vec{\theta}$ (*i.e.*, $\mathbb{E}[\vec{\mathbf{T}}(X)] = \vec{\mathbf{g}}(\vec{\theta})$), and $\mathcal{I}(\vec{\theta})$ is the expected Fisher information matrix, then



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$$\operatorname{Var}[\vec{\mathbf{T}}(X)] \geq \left(\frac{\partial \vec{\mathbf{g}}(\vec{\theta})}{\partial \vec{\theta}}\right) \mathcal{I}^{-1}(\vec{\theta}) \left(\frac{\partial \vec{\mathbf{g}}(\vec{\theta})}{\partial \vec{\theta}}\right)^{\mathrm{T}}$$

Cramér-Rao Lower Bound (CRLB)



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Cramér-Rao Lower Bound (CRLB)

In particular, if we set $g(\vec{\theta}) = \vec{\theta}$, so that $\vec{\mathsf{T}}(X)$ is an unbiased estimator for $\vec{\theta}$,

 $\operatorname{Var}[\vec{\mathsf{T}}(X)] \geq \mathcal{I}^{-1}(\vec{\theta})$

The inverse of the Fisher Information (\equiv covariance) of a parameter is a lower bound on the variance of any unbiased estimator of that parameter.



If $\vec{\mathbf{T}}(X)$ is an unbiased estimator of a function $\vec{\mathbf{g}}(\vec{\theta})$ of the parameters $\vec{\theta}$ (*i.e.*, $\mathbb{E}[\vec{\mathbf{T}}(X)] = \vec{\mathbf{g}}(\vec{\theta})$), and $\mathcal{I}(\vec{\theta})$ is the expected Fisher information matrix, then

$$\operatorname{Var}[\vec{\mathbf{T}}(X)] \geq \left(\frac{\partial \vec{\mathbf{g}}(\vec{\theta})}{\partial \vec{\theta}}\right) \mathcal{I}^{-1}(\vec{\theta}) \left(\frac{\partial \vec{\mathbf{g}}(\vec{\theta})}{\partial \vec{\theta}}\right)^{\mathrm{T}}$$

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For biased estimators: If $\mathbb{E}[\vec{T}(X) - \vec{\theta}] = \vec{B}(\vec{\theta}) \neq 0$, set $\vec{g}(\vec{\theta}) = \vec{B}(\vec{\theta}) + \vec{\theta}$ and apply CRLB.



Recall:

$$\frac{\partial \ell}{\partial \mu} = \sum_{i=1}^{N} \left(\frac{x_i - \mu}{\sigma^2} \right) \qquad \frac{\partial \ell}{\partial \sigma} = \frac{1}{\sigma} \left(-N + \sum_{i=1}^{N} \frac{(x_i - \mu)^2}{\sigma^2} \right) \qquad \mathbb{E} \Big[\sum_{i=1}^{N} (x_i - \mu)^2 \Big] = N \ \sigma^2$$

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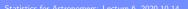
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Log-likelihood is fed to routine by user.

Routine optimises this function using a variety of techniques.

The output will include the MLE as well as the covariance matrix.

Example: fitting a line to data with uncertainties.



"It is worth checking any piece of remembered statistics, as it is almost certain to be based on the Gaussian distribution."

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Point estimation requires two decisions:



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Model specification: required to compute the likelihood. How do we know it is correct? Model validation (goodness-of-fit). Model selection.



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Point estimation requires two decisions:

Model specification: required to compute the likelihood. How do we know it is correct? Model validation (goodness-of-fit). Model selection.

Estimation method: which estimator do we pick?

The MLE is not always unbiased.

Minimum Variance Unbiased Estimator (MVUE) – among unbiased estimators, pick the one with the least variance.

