



Statistics for Astronomers: Lecture 12, 2020.11.23

Prof. Sundar Srinivasan

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Empirical distribution function.

Bootstrap. For bootstrap Cls, good discussion <u>€ here</u>.



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Hypothesis testing



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Barlow AstroML



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Ronald Fisher offered [Blanche Muriel] Bristol a cup of hot tea that he had just drawn from an urn. Bristol declined it, saying that she preferred the flavour when the milk was poured into the cup before the tea. Fisher scoffed that the order of pouring could not affect the flavour. Bristol insisted that it did and that she could tell the difference. Overhearing this debate, William Roach said, 'Let's test her'.

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 $P(\geq 3 \text{ of } 4 \text{ cups correct by chance}): (16 + 1)/70 \approx 23\%.$ $P(4 \text{ of } 4 \text{ cups correct by chance}): 1/70 \approx 1.4\% < 5\%.$



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Bristol correctly characterised all eight cups.





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What kinds of questions can be answered? (from Barlow)				
What is the straight line fit for y vs. x ?	Does y increase with x?			
What is the strength of the effect?	<i>Is</i> the effect present?			
What are the values of a and b?	Do a and b have the same value?			
Formulate the question precisely by expressing it as a hypothesis.				



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Statistical test: Procedure. Input: samples. Computes: test statistic. Output: a hypothesis.

Hypothesis: assertion/statement that can be tested using observations (e.g., "the population mean is < 5").



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Can be two-tailed/non-directional test e.g., " $\theta = \theta_0$ ", " $-5 \le \mu \le 5$ ". or one-tailed/directional test e.g., " $\theta > \theta_0$ ", " $\mu < 5$ ".

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Typically, a statement expressing lack of correlation between observations and the suggested model (*i.e.*, the data are not significantly different from noise), and the alternate hypothesis H_A suggests a relationship.

Want to demonstrate that $\langle effect \rangle exists$? Start by stating it doesn't, then find out whether data provides enough evidence to reject H_0 – hypothesis testing.



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W. H. Self, et al. JAMA,

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Frequentist inference: probability that a given hypothesis is correct is either 0 or 1. Just because we reject H_0 on the basis of <u>one dataset</u> doesn't mean H_0 is wrong or H_A is correct. At 95% confidence, frequentist procedure will reject H_0 for 5% of datasets drawn from H_0 !





Choose null (H_0) and alternate $(H_A \text{ or } H_1)$ hypotheses. Compute significance (α) using data. α is the level of tolerance for incorrectly rejecting H_0 . Outcome "significant" if small chance of occurrence from H_0 . Given α , only two possible outcomes: reject/unable to reject H_0 .



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		Reject	Type I error (false positive) (probability = a)	Correct inference (true positive) (probability = 1-β)	Source:



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Rejection or critical region: set of test statistic values for which we are able to reject H_0 .

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Critical value: the threshold separating acceptance and rejection regions.



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Error rates:

Type I (false +ve, false alarm): $P(\text{reject } H_0 \mid H_0 \text{ true}) \equiv \alpha$. Type II (false -ve): $P(\text{don't reject } H_0 \mid H_0 \text{ false}) \equiv \beta$.



Type I/II errors and classification algorithms

Classification typically involves placing boundaries in multidimensional parameter space to separate "clusters" of objects.



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Example: YSO researcher wants to identify promising massive embedded YSO candidates for spectroscopic follow-up. She devises cuts in colour-magnitude and colour-colour space to separate "high-reliability" YSO candidates from other kinds of sources with similar colours (*e.g.*, highly evolved dusty AGB stars, background galaxies).



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Type I error = false +ve = contamination ("spurious detections") to the YSO candidate sample. Type II error = false -ve rate ("missed sources") reduces the completeness of the YSO candidate sample. Compromise between increasing completeness and decreasing contamination - received operating characteristic (ROC) curve (true +ve rate vs. true -ve rate).

See Sec. 4.6.1 in the AstroML book.



Hypothesis testing: basic procedure

- Identify a null hypothesis and an alternate hypothesis, choose significance threshold lpha.
- 2 Design test statistic T. Assuming H_0 is true, obtain the distribution of T.
 - Usually complicated/unknown; use the asymptotic distribution $(N \rightarrow \infty)$.
- I Using the data, compute t, the observed value of T.
- Ompute the *p*-value: $p \equiv P(T = t | H_0 \text{ is true})$.
- If the $p < \alpha$, the tolerance for false negatives, reject H_0 at significance level α .



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Example 1

<u>Observation</u>: Tossing a coin 10 times, we observe 9 heads. <u>Statistic</u>: S_{10} , the total number of heads in 10 tosses. <u> H_0 </u>: fair coin. Under H_0 , $S_{10} \sim \text{Binomial}(1/2)$. <u> H_A </u>: $p \neq 0.5$ (two-tailed). Significance chosen: $\alpha = 0.05$.



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Since the *p*-value (= 0.009) < significance, reject H_0 at significance level $\alpha = 0.05$.



Example 2 (Barlow 8.2.2)

55% of patients suffering from a disease are spontaneously cured within a week. A new medication is tested on 105 patients. How many patients need to be cured in a week to decide whether the medication is effective at 5% significance?

 $H_0: p \le 0.55; H_A: p > 0.55$ (one-tailed test)

Statistic: k, the total number of people cured within a week.

 $k \sim \text{Binomial}(0.55)$ under null hypothesis.

Significance chosen: $\alpha = 0.05$.

We are looking for k_{α} such that $P(k \ge k_{\alpha}) < \alpha$.

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"To reject H_0 at 5% significance, more than () patients need to be cured within a week."





Statistical power of a test: $P(\text{reject } H_0 \mid H_0 \text{ false}) \equiv 1 - \beta$ As critical/threshold value \uparrow , $\alpha \downarrow$ but power also \downarrow .

Efficiency of a test: sample size required to achieve a given power.



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If both H_0 and H_A are simple, $p_{\tau}(t \mid H_0 \text{ true})$ and $p_{\tau}(t \mid H_A \text{ true})$ known. \implies the likelihood ratio is the most powerful test statistic.





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 $\text{Likelihood ratio} = \frac{\text{likelihood } H_{\text{A}} \text{ true given data}}{\text{likelihood } H_{0} \text{ true given data}} > \text{threshold} \Longrightarrow \text{reject } H_{0}.$





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Likelihood ratio = $\frac{\text{likelihood } H_A \text{ true given data}}{\text{likelihood } H_0 \text{ true given data}} > \text{threshold} \implies \text{reject } H_0.$ If H_0 , H_A simple, write in terms of parameter values: $LR = \frac{\mathscr{L}(\theta = \theta_1 \mid H_A \text{ true})}{\mathscr{L}(\theta = \theta_0 \mid H_0 \text{ true})} > \text{threshold}.$ The value of the threshold is picked such that the false-alarm probability is α .





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Typically, for convenience, written in terms of log-likelihood. Recall: for Gaussian variables, $\ln \mathscr{L} = \mathrm{constant} - \frac{1}{2}\chi^2$. Wilks' Theorem: asymptotic behavior of $\ln LR$ under H_0 is χ^2 !



 $X_i(i=1,\cdots,N)\sim \mathscr{N}(\mu,\sigma^2)$ with $\sigma=4$ and μ unknown. H_0 : $\mu=0$, H_A : $\mu=4$.

Find N and LR threshold such that we are able to reject H_0 at significance lpha=0.05 and our test has power 1-eta=0.95.



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Find N and LR threshold such that we are able to reject H_0 at significance $\alpha = 0.05$ and our test has power $1 - \beta = 0.95$.

For both hypotheses,
$$\mathscr{L}(\mu) = \prod_{i=1}^{N} \left(\frac{1}{\sqrt{2\pi}\sigma}\right)^{N} \exp\left[-\frac{1}{2}\left(\frac{x_{i}-\mu}{\sigma}\right)^{2}\right] \Longrightarrow \ln \mathscr{L}(\mu) = \text{const.} - \frac{1}{2\sigma^{2}}\sum_{i=1}^{N} (x_{i}-\mu)^{2}$$



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Recall: CLT means that $\overline{x} \sim \mathcal{N}(\mu, \sigma^2/N)$.



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Do they have the same variance? Use the *F*-test.





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See documentation for statsmodels.stats.weightstats.ztest - options and alternatives!

Now, assume N = 20. We have to use the *t*-statistic.

Can the inspector reject the company's claim at the 5% level?

See documentation for scipy.stats.ttest_1samp - options and alternatives!

