



Statistics for Astronomers: Lecture 14, 2020.12.04

Prof. Sundar Srinivasan

IRyA/UNAM





Prof. Sundar Srinivasan - IRyA/UNAM

Hypothesis testing.

Two-sample Z- and t-tests.



Prof. Sundar Srinivasan - IRyA/UNAM

・ロト・日・・日・・日・ うへの



Download this Jupyter notebook.

(2) Navigate to Colaboratory.

Sign in

Click on "Upload" and upload the notebook you downloaded in step 1.



Prof. Sundar Srinivasan - IRyA/UNAM

Data: two independent samples of sizes N_1, N_2 drawn from $\mathcal{N}(\mu_1, \sigma_1^2)$ and $\mathcal{N}(\mu_2, \sigma_2^2)$. Question: Is $\sigma_1 \neq \sigma_2$? Null hypothesis H_0 : $\sigma_1 = \sigma_2$.



Prof. Sundar Srinivasan - IRyA/UNAM

Data: two independent samples of sizes N_1 , N_2 drawn from $\mathcal{N}(\mu_1, \sigma_1^2)$ and $\mathcal{N}(\mu_2, \sigma_2^2)$. Question: Is $\sigma_1 \neq \sigma_2$? Null hypothesis H_0 : $\sigma_1 = \sigma_2$.

Recall: If $X \sim \mathcal{N}(\mu, \sigma^2)$, the sample variance $S^2 \sim \sigma^2 \frac{\chi^2(N-1)}{N-1} = \sigma^2 \chi^2_{\text{red}}(N-1)$.



Prof. Sundar Srinivasan - IRyA/UNAM

Data: two independent samples of sizes N_1, N_2 drawn from $\mathscr{N}(\mu_1, \sigma_1^2)$ and $\mathscr{N}(\mu_2, \sigma_2^2)$. Question: Is $\sigma_1 \neq \sigma_2$? Null hypothesis H_0 : $\sigma_1 = \sigma_2$.

 $\begin{array}{ll} \text{Recall: If } X \sim \mathscr{N}(\mu, \sigma^2), \text{ the sample variance } S^2 \sim \sigma^2 \ \frac{\chi^2(N-1)}{N-1} = \sigma^2 \chi^2_{\mathrm{red}}(N-1). \\ \\ \text{Let } S_2^2 > S_1^2. \text{ Define the ratio } F = \frac{S_2^2}{S_1^2} = \frac{\sigma_2^2}{\sigma_1^2} \ \frac{\chi^2_{\mathrm{red}}(N_2-1)}{\chi^2_{\mathrm{red}}(N_1-1)}. \\ \\ \text{Under } H_0, \ F = \frac{\chi^2_{\mathrm{red}}(N_2-1)}{\chi^2_{\mathrm{red}}(N_1-1)} \quad \text{ by design, } F > 1. \end{array}$



Prof. Sundar Srinivasan - IRyA/UNAM

Data: two independent samples of sizes N_1 , N_2 drawn from $\mathcal{N}(\mu_1, \sigma_1^2)$ and $\mathcal{N}(\mu_2, \sigma_2^2)$. Question: Is $\sigma_1 \neq \sigma_2$? Null hypothesis H_0 : $\sigma_1 = \sigma_2$.

Recall: If $X \sim \mathcal{N}(\mu, \sigma^2)$, the sample variance $S^2 \sim \sigma^2 \frac{\chi^2(N-1)}{N-1} = \sigma^2 \chi^2_{red}(N-1)$. Let $S_2^2 > S_1^2$. Define the ratio $F = \frac{S_2^2}{S^2} = \frac{\sigma_2^2}{\sigma^2} \frac{\chi^2_{red}(N_2-1)}{\chi^2(N_2-1)}$.

Under
$$H_0$$
, $F = rac{\chi^2_{
m red}(N_2-1)}{\chi^2_{
m red}(N_1-1)}$ by design, $F > 1$.

This quantity, the ratio of two reduced χ^2 variables, has an F distribution with parameters ($\nu_{numerator}, \nu_{denominator}$) = ($N_2 - 1, N_1 - 1$).

Python implementation: scipy.stats.f. Also see Sec. 4.7.6 in the AstroML book.



Data: two independent samples of sizes N_1 , N_2 drawn from $\mathcal{N}(\mu_1, \sigma_1^2)$ and $\mathcal{N}(\mu_2, \sigma_2^2)$. Question: Is $\sigma_1 \neq \sigma_2$? Null hypothesis H_0 : $\sigma_1 = \sigma_2$.

Recall: If $X \sim \mathcal{N}(\mu, \sigma^2)$, the sample variance $S^2 \sim \sigma^2 \frac{\chi^2(N-1)}{N-1} = \sigma^2 \chi^2_{red}(N-1)$.

Let
$$S_2^2 > S_1^2$$
. Define the ratio $F = \frac{S_2^2}{S_1^2} = \frac{\sigma_2^2}{\sigma_1^2} \frac{\chi_{red}(N_2 - 1)}{\chi_{red}^2(N_1 - 1)}$.
Under H_0 , $F = \frac{\chi_{red}^2(N_2 - 1)}{\chi_{red}^2(N_1 - 1)}$ by design, $F > 1$.

This quantity, the ratio of two reduced χ^2 variables, has an F distribution with parameters ($\nu_{numerator}, \nu_{denominator}$) = ($N_2 - 1, N_1 - 1$).

Python implementation: scipy.stats.f. Also see Sec. 4.7.6 in the AstroML book.

Example (from Barlow): see Jupyter notebook.



Data: two independent samples of sizes N_1 , N_2 drawn from $\mathcal{N}(\mu_1, \sigma_1^2)$ and $\mathcal{N}(\mu_2, \sigma_2^2)$. Question: Is $\sigma_1 \neq \sigma_2$? Null hypothesis H_0 : $\sigma_1 = \sigma_2$.

Recall: If $X \sim \mathscr{N}(\mu, \sigma^2)$, the sample variance $S^2 \sim \sigma^2 \; rac{\chi^2(N-1)}{N-1} = \sigma^2 \chi^2_{
m red}(N-1)$.

Let
$$S_2^2 > S_1^2$$
. Define the ratio $F = \frac{S_2^2}{S_1^2} = \frac{\sigma_2^2}{\sigma_1^2} \frac{\chi_{\rm red}^2(N_2 - 1)}{\chi_{\rm red}^2(N_1 - 1)}$
Under H_0 , $F = \frac{\chi_{\rm red}^2(N_2 - 1)}{\chi_{\rm red}^2(N_1 - 1)}$ by design, $F > 1$.

This quantity, the ratio of two reduced χ^2 variables, has an *F* distribution with parameters ($\nu_{numerator}, \nu_{denominator}$) = ($N_2 - 1, N_1 - 1$).

Python implementation: scipy.stats.f. Also see Sec. 4.7.6 in the AstroML book.

Example (from Barlow): see Jupyter notebook.

For you to solve:

Given
$$N_1=10, N_2=7$$
, and $S_2^2=(1+\lambda)S_1^2$, with $\lambda>0$.

What is λ if the *F*-test rejects the null hypothesis with 99.5% confidence?



Application of F-test: detecting time-variability

Data: fluxes f_i ($i = 1, \dots, N$) for a source observed at times t_i . Question: Are the data consistent with a variable source?



Prof. Sundar Srinivasan - IRyA/UNAM

Data: fluxes f_i $(i = 1, \dots, N)$ for a source observed at times t_i .

Question: Are the data consistent with a variable source?

- H_0 : No. Variability solely due to measurement error.
- H_A : Observed time dependence is combination of intrinsic variability and measurement error.



Data: fluxes f_i $(i = 1, \dots, N)$ for a source observed at times t_i .

Question: Are the data consistent with a variable source?

 H_0 : No. Variability solely due to measurement error.

 H_{a} : Observed time dependence is combination of intrinsic variability and measurement error.



Data: fluxes f_i $(i = 1, \dots, N)$ for a source observed at times t_i .

Question: Are the data consistent with a variable source?

 H_0 : No. Variability solely due to measurement error.

 H_{A} : Observed time dependence is combination of intrinsic variability and measurement error.

Model:
$$f_i = f_{\text{true},i} + f_{\text{meas},i}$$
, with $f_{\text{meas},i} \sim \mathcal{N}(0, \sigma^2)$.
 $\operatorname{Var}[f_i] = \operatorname{Var}[f_{\text{true},i}] + \operatorname{Var}[f_{\text{meas},i}] = \operatorname{Var}[f_{\text{true},i}] + \sigma^2$.
Under H_0 , $\operatorname{Var}[f_i] = \sigma^2$. Under H_A , $\operatorname{Var}[f_i] > \sigma^2$.

Perform *F*-test on the data! Define $F = \frac{\text{Sample variance for data } f_i}{\text{Sample variance of measurement error}}$

See Jupyter notebook for illustration.





Statistics for Astronomers: Lecture 14, 2020.12.04

Prof. Sundar Srinivasan - IRyA/UNAN

▶ This series of videos-Barlow Wall & Jenkins AstroML



Prof. Sundar Srinivasan - IRyA/UNAM

・ロット 4回ッ 4回ッ 4回ッ 4回ッ

How they work: typically by sorting/ordering (ranking) data values.



Prof. Sundar Srinivasan - IRyA/UNAM

How they work: typically by sorting/ordering (ranking) data values.

Some advantages:



 Very little assumed about the data and underlying distributions (typically: continuity). Great for when we don't know!

- Ideal for small sample sizes, poorly-defined distributions, data with outliers. (2) difficult to determine parametric model for such data.
- (3) Also applicable when data is non-numeric (*e.g.*, classifications).
 - Can work with samples drawn from several different populations.

Major drawback: some tests require binning of data.



How they work: typically by sorting/ordering (ranking) data values.

Some advantages:



Very little assumed about the data and underlying distributions (typically: continuity). Great for when we don't know!

- Ideal for small sample sizes, poorly-defined distributions, data with outliers. difficult to determine parametric model for such data.
- (3) Also applicable when data is non-numeric (*e.g.*, classifications).
 - Can work with samples drawn from several different populations.

Major drawback: some tests require binning of data.

We've already used nonparametric inference!

Location estimate: sort data, pick central value - median. Scale estimate: sort data, compute interguartile range. Distribution estimate: EDF. Used this for interval estimates (CI).

The bootstrap procedure.





One-sample test (implementation: scipy.stats.kstest):

Given: sample of size *N*. Question: Is the sample drawn from a particular distribution? H_0 : yes. H_A : No (two-sided) or greater/less than CDF of distribution (one-sided).

Two-sample test (implementation: scipy.stats.ks_2samp):

Given: samples of sizes N_1 , N_2 . Question: Are samples drawn from the same distribution? H_0 : yes. H_A : No (two-sided).



One-sample test (implementation: scipy.stats.kstest):

Given: sample of size *N*. Question: Is the sample drawn from a particular distribution? H_0 : yes. H_A : No (two-sided) or greater/less than CDF of distribution (one-sided).

Two-sample test (implementation: scipy.stats.ks_2samp):

Given: samples of sizes N_1 , N_2 . Question: Are samples drawn from the same distribution? H_0 : yes. H_A : No (two-sided).

KS statistic: maximum distance between CDFs that are being compared. One-sample case: $D_{KS} = \max \left| CDF_{model}(x) - ECDF(x) \right|$ (two-sided). $D_{KS} = \max \left(CDF_{model}(x) - ECDF(x) \right)$ (one-sided).



of. Sundar Srinivasan - IRyA/UNAM

One-sample test (implementation: scipy.stats.kstest):

Given: sample of size *N*. Question: Is the sample drawn from a particular distribution? H_0 : yes. H_A : No (two-sided) or greater/less than CDF of distribution (one-sided).

Two-sample test (implementation: scipy.stats.ks_2samp):

Given: samples of sizes N_1 , N_2 . Question: Are samples drawn from the same distribution? H_0 : yes. H_A : No (two-sided).

KS statistic: maximum distance between CDFs that are being compared. One-sample case: $D_{KS} = \max \left| CDF_{model}(x) - ECDF(x) \right|$ (two-sided). $D_{KS} = \max \left(CDF_{model}(x) - ECDF(x) \right)$ (one-sided).

Let's go to the Jupyter notebook again...



One-sample test (implementation: scipy.stats.kstest):

Given: sample of size *N*. Question: Is the sample drawn from a particular distribution? H_0 : yes. H_A : No (two-sided) or greater/less than CDF of distribution (one-sided).

Two-sample test (implementation: scipy.stats.ks_2samp):

Given: samples of sizes N_1 , N_2 . Question: Are samples drawn from the same distribution? H_0 : yes. H_A : No (two-sided).

KS statistic: maximum distance between CDFs that are being compared. One-sample case: $D_{\rm KS} = \max \left| CDF_{\rm model}(x) - ECDF(x) \right|$ (two-sided). $D_{\rm KS} = \max \left(CDF_{\rm model}(x) - ECDF(x) \right)$ (one-sided).

Let's go to the Jupyter notebook again...

Advantages: No binning! Small samples! More powerful for intermediate-size samples! Can also work as a one-tailed test (see sscipy.stats.kstest).

Disadvantages: Not sensitive to differences in the tails. Doesn't care about #dof.



One-sample test (implementation: scipy.stats.kstest):

Given: sample of size *N*. Question: Is the sample drawn from a particular distribution? H_0 : yes. H_A : No (two-sided) or greater/less than CDF of distribution (one-sided).

Two-sample test (implementation: scipy.stats.ks_2samp):

Given: samples of sizes N_1 , N_2 . Question: Are samples drawn from the same distribution? H_0 : yes. H_A : No (two-sided).

KS statistic: maximum distance between CDFs that are being compared. One-sample case: $D_{\rm KS} = \max \left| CDF_{\rm model}(x) - ECDF(x) \right|$ (two-sided). $D_{\rm KS} = \max \left(CDF_{\rm model}(x) - ECDF(x) \right)$ (one-sided).

Let's go to the Jupyter notebook again...

Advantages: No binning! Small samples! More powerful for intermediate-size samples! Can also work as a one-tailed test (see sscipy.stats.kstest).

Disadvantages: Not sensitive to differences in the tails. Doesn't care about #dof.

